

What about nonconjugate priors?

(1)

$$p(\theta | \vec{y}) \propto \boxed{\vec{y} = y_1, \dots, y_n}$$

$$\prod_{i=1}^n \theta^{y_i} \frac{e^{-\theta}}{y_i!}$$

$Y|\theta \sim \text{Poisson}$
 $\theta > 0$

likelihood
 $p(y_1, \dots, y_n | \theta)$

$$\frac{1}{\theta \sigma \sqrt{2\pi}} \exp \left\{ \frac{-(\ln \theta - \mu)^2}{2\sigma^2} \right\}$$

Prior (log-normal)
 $p(\theta)$

$$\mathbb{E} \theta | y_1, \dots, y_n = ?$$

Exercise 1

$$p(\theta) = \frac{1}{2}$$

$$\boxed{0 < \theta < 2}$$

$$\phi = \log \theta \Rightarrow \theta = e^\phi$$

want $p(\phi)$

$$\rightarrow p(\phi) d\phi = p(\theta) d\theta$$

$$p(\phi) = \underbrace{p(\theta)}_{\frac{1}{2}} \underbrace{\left| \frac{d\theta}{d\phi} \right|}_{e^\phi}$$

$$p(\phi) = \frac{1}{2} e^\phi$$

$$-\infty < \phi < \log 2$$

Exercise 2

2

We want $\vec{y}_2 = \{y_{21}, \dots, y_{2n_2}\}$

$$P(\tilde{y}_1 < \tilde{y}_2 | \vec{y}_1, \vec{y}_2)$$

new obs
of #
children
to a
woman
w/o
bachelors

same
but
"w/
bachelors"

$$\vec{y}_1 = \{y_{11}, \dots, y_{1n_1}\}$$

To compute w/ Monte Carlo sampling
we need to be able to sample from

$$P(\tilde{y}_1 | \vec{y}_1, \vec{y}_2) \text{ and}$$

$$P(\tilde{y}_2 | \vec{y}_1, \vec{y}_2)$$

By independence assumed in our model:

$$P(\tilde{y}_1 | \vec{y}_1, \vec{y}_2) = P(\tilde{y}_1 | \vec{y}_1)$$

$$P(\tilde{y}_2 | \vec{y}_1, \vec{y}_2) = P(\tilde{y}_2 | \vec{y}_2)$$

In general, when $\tilde{y}$ is a new obs.
from the same population, we call

$$P(\tilde{y} | \vec{y})$$

the

"posterior
predictive
distribution"

$$\vec{y}_1 \rightarrow \theta_1 \rightarrow \tilde{y}_1$$

$$\vec{y}_2 \rightarrow \theta_2 \rightarrow \tilde{y}_2$$